# Providence [Providence Digital Commons](https://digitalcommons.providence.org/)

[Articles, Abstracts, and Reports](https://digitalcommons.providence.org/publications) 

5-1-2017

# Sex-specific associations of maternal birthweight with offspring birthweight in the Omega study.

Collette N Ncube

Amelia R Gavin

Michelle A Williams Center for Perinatal Studies, Swedish Medical Center, Seattle, WA, USA

Chunfang Qiu Center for Perinatal Studies, Swedish Medical Center, Seattle, WA

Tanya K Sorensen Center for Perinatal Studies, Swedish Medical Center, Seattle, WA

See next page for additional authors

Follow this and additional works at: [https://digitalcommons.providence.org/publications](https://digitalcommons.providence.org/publications?utm_source=digitalcommons.providence.org%2Fpublications%2F1344&utm_medium=PDF&utm_campaign=PDFCoverPages)

Part of the [Oncology Commons,](https://network.bepress.com/hgg/discipline/694?utm_source=digitalcommons.providence.org%2Fpublications%2F1344&utm_medium=PDF&utm_campaign=PDFCoverPages) and the [Other Medical Specialties Commons](https://network.bepress.com/hgg/discipline/708?utm_source=digitalcommons.providence.org%2Fpublications%2F1344&utm_medium=PDF&utm_campaign=PDFCoverPages)

## Recommended Citation

Ncube, Collette N; Gavin, Amelia R; Williams, Michelle A; Qiu, Chunfang; Sorensen, Tanya K; and Enquobahrie, Daniel A, "Sex-specific associations of maternal birthweight with offspring birthweight in the Omega study." (2017). Articles, Abstracts, and Reports. 1344. [https://digitalcommons.providence.org/publications/1344](https://digitalcommons.providence.org/publications/1344?utm_source=digitalcommons.providence.org%2Fpublications%2F1344&utm_medium=PDF&utm_campaign=PDFCoverPages)

This Article is brought to you for free and open access by Providence Digital Commons. It has been accepted for inclusion in Articles, Abstracts, and Reports by an authorized administrator of Providence Digital Commons. For more information, please contact [digitalcommons@providence.org.](mailto:digitalcommons@providence.org)

## Authors

Collette N Ncube, Amelia R Gavin, Michelle A Williams, Chunfang Qiu, Tanya K Sorensen, and Daniel A Enquobahrie

This article is available at Providence Digital Commons:<https://digitalcommons.providence.org/publications/1344>



# **HHS Public Access**

Author manuscript Ann Epidemiol. Author manuscript; available in PMC 2018 May 10.

Published in final edited form as:

Ann Epidemiol. 2017 May ; 27(5): 308–314.e4. doi:10.1016/j.annepidem.2017.04.006.

# **Sex-specific Associations of Maternal Birthweight with Offspring Birthweight in the Omega Study**

**Collette N. Ncube**a, **Amelia R. Gavin**b, **Michelle A. Williams**c,d, **Chunfang Qiu**<sup>c</sup> , **Tanya K.**  Sorensen<sup>c</sup>, and Daniel A. Enquobahrie<sup>a,c</sup>

<sup>a</sup>Department of Epidemiology, School of Public Health, University of Washington, 1959 NE Pacific Street, Health Sciences Building, Seattle, WA 98195-7236, USA

<sup>b</sup>School of Social Work, University of Washington, 4101 15th Ave NE, Seattle, WA 98105, USA

<sup>c</sup>Center for Perinatal Studies, Swedish Medical Center, 721 Minor Avenue, Seattle, WA 98104, USA

<sup>d</sup>Department of Epidemiology, Harvard T.H. Chan School of Public Health, 677 Huntington Avenue, Boston, MA 02115, USA

# **Abstract**

**Purpose—**We investigated nonlinear and offspring-sex specific associations of maternal birthweight (BW) with offspring BW among participants of the Omega study, a pregnancy cohort.

**Methods—**Maternal BW was modeled as a continuous variable, linear spline, and binary variable indicating low birthweight (LBW) (<2500 vs. ≥2500grams). Offspring BW was modeled as a continuous and binary variable in regression models. Non-linearity was assessed using likelihood ratio tests (LRT) in marginal linear spline models.

**Results—**For every 100gram increase of maternal BW, offspring BW increased by 22.29 (95%CI: 17.57, 27.02) or 23.41 (95%CI: 6.87, 39.96) grams among mothers with normal BW or born macrosomic, respectively, but not among LBW mothers (β=–8.61 grams; 95%CI: –22.88, 5.65) (LRT p-value=0.0005). For every 100gram increase in maternal BW, BW of male offspring increased 23.47 (95%CI: 16.75, 30.19) or 25.21 (95%CI: 4.35, 46.07) grams among mothers with normal BW or born macrosomic, respectively, while it decreased 31.39 grams (95%CI: −51.63, −11.15) among LBW mothers (LRT p-value<0.0001). Corresponding increases in BW of female offspring (16–22 grams) did not differ among mothers with LBW, normal BW or macrosomia (LRT p-value=0.9163).

**Conclusions—**Maternal and offspring BW associations are evident among normal BW and macrosomic mothers. These associations differ by offspring sex.

**Correspondence**: Collette N. Ncube, University of Washington, Box 357236, 1959 NE Pacific Street, Seattle, WA 98195-7236 USA, ncubec@u.washington.edu, Telephone: +1 (206) 543-7559.

**Publisher's Disclaimer:** This is a PDF file of an unedited manuscript that has been accepted for publication. As a service to our customers we are providing this early version of the manuscript. The manuscript will undergo copyediting, typesetting, and review of the resulting proof before it is published in its final citable form. Please note that during the production process errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.

#### **Keywords**

Sex; birthweight; infant; low birthweight; mothers; body mass index; weight gain; pregnancy

## **INTRODUCTION**

Birthweight (BW) is an indicator of fetal growth and development [1] which are important determinants of life course health. Low birthweight (LBW), less than 2500 grams, is associated with an increase in risk for morbidity and mortality in infancy [2, 3], and chronic diseases in adulthood [4–6]. LBW has a multifactorial origin [7]. Several proximal risk factors including those during or immediately prior to the pregnancy (e.g., maternal age and pre-pregnancy body mass index (ppBMI)) have been identified [7, 8]. From a life course perspective, distal risk factors such as mothers' BW, childhood health, and early life socioeconomic position affect later life pregnancy outcomes [9]. These distal risk factors may be influential in the perpetuation of poor birth outcomes among certain groups.

Ounsted and Ounsted (1968) theorized that women who had constrained, in utero growth were more likely to have offspring with intrauterine growth retardation [10]. Since this seminal paper, several studies that examined maternal and offspring birth outcomes have been published [11–13]. Maternal BW has been consistently shown to be one of the strongest predictors of offspring BW [14]. Each 100 gram increase in maternal BW was associated with, on average, an additional 11–28 gram increase in offspring BW [15–18]; mothers who were LBW at their own birth had a two-fold increase in risk of having a LBW infant [19]. However, there is limited consensus concerning the potential non-linear relationships of maternal and offspring BW [20, 21] and whether the relationships differ for male and female offspring [22]. Despite the association of BW with adult BMI [23, 24] and the importance of ppBMI on the course and outcomes of the pregnancy [25], the role of maternal ppBMI as moderator of maternal-offspring BW associations has also not been examined. To address these limitations, we used a well-characterized pregnancy cohort to investigate overall and sex-specific associations between maternal and offspring BW.

# **MATERIALS AND METHODS**

#### **Study setting and study population**

The study was conducted among participants of the Omega study, a prospective cohort study (1996–2008) of pregnant women designed to examine risk factors for pregnancy complications and adverse outcomes [26]. Women were recruited from prenatal care clinics affiliated with Swedish Medical Center in Washington State, and were eligible to enroll if they were at least 18 years of age, able to speak and read English, initiated prenatal care before 16 weeks of gestation, and planned to carry the pregnancy to term and deliver at one of the two study hospitals. A total of 4602 women were enrolled in the study and 4343 had singleton live-births. We had complete BW data (for the mother and the singleton live-born offspring) for N=3804 Omega study participants. In the current analyses, we included infants with BW at least 300 grams (N=3800). Participants were then excluded from analyses if they were missing data on gestational age at delivery  $(n=2)$ , offspring sex  $(n=3)$ ,

smoking history ( $n=4$ ), gestational diabetes ( $n=48$ ), preeclampsia ( $n=1$ ), or weight gain during pregnancy (n=8). These were not mutually exclusive. The final sample for analyses included 3736 mother-offspring dyads. The protocol used in the Omega study was approved by the Institutional Review Boards of Swedish Medical Center and Tacoma General Hospital and all women provided written informed consent.

#### **Data collection**

In-person interviews by trained study personnel were conducted using structured questionnaires shortly after enrollment, on average 15.6 weeks gestation (SD=2.9 weeks). The interviews were used to collect data on socio-demographic characteristics, medical and family history of participants, including self-reported mothers' BW at their own birth in pounds and ounces, race, education, height, pre-pregnancy weight (immediately prior to the study pregnancy), age, prenatal cigarette smoking, and alcohol consumption. Pregnant women were followed until delivery. Information on infant BW in grams, gestational age at birth, offspring sex (male/female), and maternal weight within four weeks of delivery were abstracted from the hospital record after delivery, as was information on maternal health during the pregnancy and pregnancy complications.

#### **Exposure and outcome**

The primary exposure of interest was maternal BW, which was converted from pounds and ounces to grams. Maternal BW was modeled as 1) a continuous variable with each 1-unit change corresponding to a 100 gram change, 2) a linear spline with knots at 2500 grams (LBW) and 4000 grams (macrosomia), and 3) a binary variable indicating LBW status (<2500 vs. ≥2500 grams). The outcomes were offspring BW (as a continuous variable) and offspring LBW status.

#### **Effect modifiers and covariates**

Offspring sex was examined as a potential effect modifier. In secondary analyses, ppBMI was also considered as a potential effect modifier. Using World Health Organization criteria, ppBMI was calculated using weight  $(kg)/[\text{height}(m)]^2$  and the following categories: underweight (<18.5 kg/m<sup>2</sup>), normal weight (18.5–24.99 kg/m<sup>2</sup>) and overweight/obese (25  $\text{kg/m}^2$ ). Race (white, black, Asian, or other), preterm birth (<37 and 37 weeks gestation), family history of diabetes (yes/no), smoking history (never, current, or former smoker), educational attainment ( $high school$ ) school), maternal age ( $\langle 25, 25-35,$  or  $>35$ years), marital status (married/unmarried), parity (nulliparous/multiparous), gestational diabetes (yes/no), preeclampsia (yes/no), weight gain during pregnancy (inadequate, adequate or excessive based on Institute of Medicine recommendations per ppBMI category) [27], and chronic hypertension (yes/no) were included as covariates in statistical analyses.

#### **Statistical analyses**

We used summary statistics, means (standard deviation) and counts (percentage) for continuous and categorical variables, respectively, to describe the study population. We examined overall maternal-offspring BW associations, fitting linear regression models to estimate beta coefficients (β) and 95% confidence intervals (CIs). Maternal BW was

modeled as a continuous variable, linear spline [28], and binary variable (based on LBW status). In the first scenario, the slope estimated the average difference in mean offspring BW associated with a 100 gram increase in maternal BW. In the second scenario, the slope estimated differences in mean offspring BW per 100 gram increase in maternal BW among LBW ( $\langle 2500 \text{ grams} \rangle$ , normal BW (2500–3999 grams), and macrosomic ( $\langle 4000 \text{ grams} \rangle$ mothers. The statistical significance of the change in slope was determined using p-values of the coefficients obtained from a marginal linear spline model. We used the likelihood ratio test (LRT) to test the hypothesis that the maternal-offspring BW relationship was linear, against the alternative that it was not linear throughout the entire distribution of maternal BW. In the third scenario, we estimated the difference in mean BW of offspring delivered by LBW mothers compared to non-LBW mothers. We fit three Models in these analyses: Model 1 (unadjusted), Model 2 (adjusted for a priori determined potential confounders and precision variables selected based upon our intergenerational conceptual model: maternal race, family history of diabetes, smoking history and educational attainment, maternal age, marital status, parity, and offspring sex), and Model 3 (adjusted for Model 2 variables and potential mediators of associations: ppBMI, preterm birth, chronic hypertension, and pregnancy complications: gestational diabetes and preeclampsia). We also fit logistic regression models to estimate the odds ratios (ORs) and corresponding 95% CIs of offspring LBW associated with maternal BW modeled as a continuous variable, linear spline, and binary variable, as described above. We examined effect modification by offspring sex by repeating the analyses stratified by offspring sex. To test the statistical significance of the interactions, we fit models with indicators for maternal BW, offspring sex, and an interaction term between maternal BW and offspring sex. The p-value of the interaction term was used to determine the statistical significance of the multiplicative interaction.

In secondary analyses, we examined effect modification by ppBMI, among male and female offspring separately, by fitting the previously described models, stratified by ppBMI (normal and overweight/obese). We also fit models with indicators for maternal BW, ppBMI, and an interaction term between maternal BW and ppBMI to determine statistical significance of the multiplicative interaction. Given the small number of women who were underweight prepregnancy (N=161), particularly in strata of both offspring sex and maternal BW (modeled as a linear spline or binary variable), this group was excluded in the ppBMI effect modification analyses.

Statistical significance was determined using a two-sided p-value<0.05. All analyses were carried out using Stata version 13.1, software (Stata Corporation, College Station, Texas).

# **RESULTS**

About half of the offspring were male (51.2%) and the majority of mothers were white (86.4%), nulliparous (62.1%), married (91.5%), and with a high school education (96.6%) (Table 1). Overall, offspring BW increased 18.51 grams (95%CI: 15.28, 21.75) and the risk of LBW decreased 5% (95%CI: 0.92, 0.98) per 100 grams increase of maternal BW (Table 2). Additional adjustment for pregnancy complications and potential mediators did not substantially alter these estimates. The increase in offspring BW, per 100 grams of maternal BW, was statistically significant among mothers with normal BW (β=22.29 grams; 95%CI:

17.57, 27.02) or macrosomia (β=23.41 grams; 95%CI: 6.87, 39.96), but not among LBW mothers ( $\beta = -8.61$  grams; 95%CI:  $-22.88$ , 5.65). The change in slope between LBW and normal BW mothers was statistically significant (p-value  $< 0.0001$ ), the change in slope between normal BW and macrosomic mothers was not statistically significant (pvalue=0.908), and the linear spline model fit the maternal-offspring BW association better than the continuous model (LRT p-value=0.0005) (Table 2).

In sex-stratified models for all offspring, regardless of maternal BW, increases in offspring BW per 100 grams of maternal BW were similar among males (β=16.90 grams; 95%CI: 12.30, 21.49) and females (β = 20.17 grams; 95%CI: 15.59, 24.74) (p-value for interaction=0.444) (Table 3). Similarly, the reduction in risk of offspring LBW per 100 grams of maternal BW was similar among male  $(OR = 0.96; 95\% CI: 0.92, 1.00)$  and female  $(OR = 0.94; 95\% CI: 0.91, 0.98)$  (p-value for interaction=0.785). Among male offspring, offspring BW increased by 23.47 grams (95%CI: 16.75, 30.19) or 25.21 grams (95%CI: 4.35, 46.07) per 100 grams of maternal BW among mothers with normal BW or macrosomia, respectively, while it decreased by 31.39 grams (95%CI: −51.63, −11.15) per 100 grams of maternal BW among LBW mothers (LRT p-value<0.0001) (Table 3). Among female offspring, offspring BW increased by 16.22 grams (95%CI: −4.12, 36.55), 20.63 grams (95%CI: 13.94, 27.31) or 21.69 grams 8 (95%CI: −6.03, 49.41) per 100 grams of maternal BW among mothers with LBW, normal BW or macrosomia (LRT p-value=0.9163). The associations observed among male and female offspring were statistically significantly different (p-value for interaction=0.0148, Figure). Female offspring of LBW mothers weighed less, on average, than female offspring of non-LBW mothers (β=−228.34; 95%CI: −313.71, −142.97) and were at increased risk of being LBW themselves (OR=2.64; 95%CI: 1.48, 4.72) (Table 3). Male offspring of LBW mothers also weighed less, on average, than male offspring of non-LBW mothers ( $\beta$ =−100.96; 95%CI: −193.19, −8.73), although the reduction of BW was not as pronounced as the reduction in BW among female offspring (pvalue for interaction=0.059). Similarly, the potential increase in risk of LBW among male offspring was not statistically significant (OR=1.16; 95%CI: 0.51, 2.61) and attenuated in comparison with the LBW risk among female offspring (p-value for interaction=0.126).

Findings from analyses stratified by ppBMI were similar, in general, to those that were observed in overall sex-stratified analyses, particularly among those with normal ppBMI (Table 4). The associations observed among male  $(p=0.8132)$  or female  $(p=0.3463)$  offspring were not statistically significantly different between normal and overweight/obese ppBMI categories.

## **DISCUSSION**

Maternal BW was positively associated with offspring BW, particularly among mothers with normal BW or macrosomia. Offspring of LBW mothers weighed less than those born to non-LBW mothers and were about twice as likely to be LBW themselves. We also found evidence for potential effect modification of the maternal-offspring BW associations by offspring sex. We identified a J-shaped relationship among males and a linear relationship among females. The reduction in offspring BW and the higher risk of offspring LBW among LBW mothers, compared with non-LBW mothers, was more pronounced and statistically

significant among female offspring. The sex-specific maternal-offspring BW associations were not modified by ppBMI.

Findings of this study are consistent with previous reports which have described an overall positive association between maternal and offspring BW [15–17] and evidence of a nonlinear relationship [18, 20, 21]. For instance, Hackman et al. and Klebanoff et al. have suggested a J-shaped relationship between mean maternal and offspring BW [20, 21]. Our findings suggest that this relationship is determined primarily by male offspring. The mechanisms by which maternal and offspring BW are associated are not fully understood. They may include shared genetic attributes and environmental exposures [29], or intergenerational socioeconomic factors and neighborhood context [30, 31] which independently influence the outcome in both mother and offspring; or fetal programming of offspring birth size, due to maternal in utero growth restriction [29]. Our study extends previous work by specifically examining maternal-offspring BW association differences across the distribution of maternal BW (i.e. LBW, normal BW, macrosomia), and supporting the conclusion that an increase in maternal BW is predictive of an increase in offspring BW only among normal BW and macrosomic mothers.

Few studies have explored offspring sex-specific differences in maternal-offspring BW associations. Carr-Hill et al. (1987) reported correlations between maternal and offspring BW among mother-daughter pairs (Pearson's correlation r = 0.219; 95%CI: 0.102, 0.330) that were similar to corresponding correlations among mother-son pairs (Pearson's correlation  $r = 0.207$ ; 95%CI: 0.082, 0.326) [22]; Voldner et al. (2009) reported similar associations from multivariable regression models for female offspring  $(\beta=184 \text{ grams per } 1$ kg of maternal BW; 95%CI: 87, 280) and male offspring (β=148 grams per 1 kg of maternal BW; 95%CI: 51, 243) [32]. To our knowledge, our study is the first to report sex-specific differences in patterns of maternal-offspring BW associations and transgenerational transmission of LBW risk. We found non-linear relationships among male offspring and linear relationships among female offspring. Maternal LBW-offspring LBW associations were more pronounced among females. The distribution of BW has been conceptualized as a Gaussian distribution with two subpopulations – a predominant normal distribution (primary component) and a residual distribution (secondary component) [33]. The births in the residual distribution are believed to be different from those in the predominant distribution, and those in the lower tail are believed to be particularly at higher risk for poor health outcomes [33]. LBW mothers are more likely to fall into the secondary component of the BW distribution. Factors that cause these births to differ from those in the primary component of the BW distribution may also modify the maternal-offspring BW association. However, additional research is needed test this hypothesis.

The role of offspring sex in associations of maternal characteristics with trajectories and ultimate potentials of fetal growth, development and adulthood health are active areas of investigation [34, 35]. Most prior research has dealt with exposures and maternal characteristics during the perinatal period [36, 37] rather than taking a life course approach. Differences in fetal growth [38] and survival [39] among male and female offspring have been well documented, although mechanisms are not well understood. Studies indicate that male and female offspring respond differently to adverse environmental exposures [40, 41]

and nutritional deficiencies [42, 43], complications of pregnancy [44] and maternal phenotypic factors [45]. The role of sex chromosomes [46] and sex-specific epigenetic programming [46, 47] in the placenta are believed to influence the functioning of the organ in a sex-specific manner, thus contributing to the sexual dimorphism of fetal growth. The sexes maximize fitness differently depending on *in utero* conditions and the timing and type of exposures or constraints. These coping strategies have implications for both fetal growth and susceptibility to disease over the life course [48]. Based on the growing literature on sex differences in fetal growth, it has been proposed that male offspring respond to the intrauterine environment so as to allow for continued normal growth, which places them at risk for compromise if exposed to subsequent insults. Female offspring, on the other hand, are believed to modify growth trajectory in order to improve chances of survival [45, 49]. The linearity and non-linearity of the maternal-offspring BW association among females and males, respectively, along with the stronger associations between maternal and offspring LBW among female offspring in our study support sexual dimorphism in the influence of maternal BW on offspring BW.

Previous studies suggest a positive association between BW and adulthood BMI [24], and several others report a positive association between ppBMI and offspring BW in offspring sex-adjusted analyses [25, 50]. To our knowledge, no prior study evaluated potential effect modification of maternal-offspring BW associations by ppBMI. In the current study, we did not find evidence for effect modification. Hyppopen et al. reported that adjustment for ppBMI in a linear regression model did not affect the maternal-offspring BW association much [51]. We conducted post-hoc analyses to examine whether ppBMI mediated maternaloffspring BW associations using the potential outcomes approach to mediation analysis [52]. Pre-pregnancy BMI mediated a small proportion of the overall maternal-offspring BW associations (3.09%; 95%CI: 2.64, 3.77; p-value=0.026), but mediation did not appear to be statistically significant in offspring sex-specific analyses (Supplementary File 3). We also conducted post-hoc analyses to evaluate potential effect modification of maternal-offspring BW associations by maternal weight gain during pregnancy. The sex-specific differences in maternal-offspring BW associations, specifically among LBW mothers, were observed only among women who had inadequate weight gain during pregnancy (Supplementary File 4).

The strengths of this study include the prospective cohort study design, the wellcharacterized study population, large sample size, the modeling of the exposure using different forms (including linear splines), examining sex-specific associations, exploring potential effect modification by ppBMI and weight gain during pregnancy, and exploring potential mediation by ppBMI. Our study also has several limitations that deserve mention. First, we used self-reported maternal BW and pre-pregnancy height and weight. This may lead to potential misclassification of the outcome and biased estimates of the association(s) of interest. However, self-reported height and weight have been found to have high sensitivity and specificity among females [53] and self-reported BW has been found to have moderate to substantial agreement with recorded BW [54, 55]. In addition, the cohort study design will minimize the risk of differential misclassification. Second, we performed complete case analyses, excluding from the final analyses participants with any missing data on the variables of interest. Almost 14% of participants with live-births were excluded through list-wise deletion, the majority of whom were missing data on the exposure of

interest. These participants were more likely to be non-white, unmarried, multiparous, obese, have lower educational attainment, and their infants were more likely to be born preterm. Complete case analysis decreases efficiency, and a violation of the untestable 'missing completely at random' assumption may lead to biased estimates. Finally, racial/ ethnic minorities were not well represented in our study population. Researchers have found potential race-specific differences in transgenerational LBW risk [56]. Unfortunately, we were not able to assess potential effect modification by race. Generalizability of our findings may be limited to other populations that have comparable characteristics to the Omega study population.

In conclusion, we found that maternal and offspring BW were positively associated, particularly among mothers with normal and macrosomic BW. Offspring sex modified maternal and offspring BW associations. Our findings highlight the importance of examining sex differences in transgenerational fetal growth studies, and, provide guidance and motivation for future investigations of potential mechanisms for maternal-offspring BW associations. This is of public health significance as it could help improve identification of populations at risk for poor birth outcomes and institute preventative and/or early diagnostic intervention.

#### **Supplementary Material**

Refer to Web version on PubMed Central for supplementary material.

# **Acknowledgments**

Authors are grateful for participants of the Omega pregnancy cohort study and staff of the Center for Perinatal Studies, Swedish Medical Center, WA.

#### **FUNDING SOURCES**

The Omega study was supported by grant from the National Institutes of Health (R01HD032562). Dr. Ncube was supported by the Reproductive, Perinatal and Pediatric Epidemiology Training Program of the National Institute of Child Health and Human Development (T32 HD052462).

## **Abbreviations**



## **References**

1. Bollen KA, Noble MD, Adair LS. Are gestational age, birth weight, and birth length indicators of favorable fetal growth conditions? A structural equation analysis of Filipino infants. Stat Med. 2013; 32(17):2950–61. DOI: 10.1002/sim.5771 [PubMed: 23494711]

- 3. Martin, JA., Hamilton, BE., Ventura, SJ., Osterman, MJK., Mathews, TJ. Births: Final Data for 2011. Hyattsville, MD: National Center for Health Statistics; 2013.
- 4. Barker DJ, Eriksson JG, Forsen T, Osmond C. Fetal origins of adult disease: strength of effects and biological basis. Int J Epidemiol. 2002; 31(6):1235–9. DOI: 10.1093/ije/31.6.1235 [PubMed: 12540728]
- 5. Ross, MG., Desai, M. Developmental Origins of Adult Health and Disease. In: Gabbe, SG.Niebyl, JR., Simpson, JL., editors. Obstetrics: Normal and Problem Pregnancies. 5th. Philadelphia (PA): Churchill Livingstone Elsevier; 2007.
- 6. Barker DJ. Fetal origins of coronary heart disease. BMJ. 1995; 311(6998):171–4. doi[:http://](http://dx.doi.org/10.1136/bmj.311.6998.171) [dx.doi.org/10.1136/bmj.311.6998.171.](http://dx.doi.org/10.1136/bmj.311.6998.171) [PubMed: 7613432]
- 7. Kramer MS. Determinants of low birth weight: methodological assessment and meta-analysis. Bull World Health Organ. 1987; 65(5):663–737. [PubMed: 3322602]
- 8. Valero De Bernabe J, Soriano T, Albaladejo R, Juarranz M, Calle ME, Martinez D, et al. Risk factors for low birth weight: a review. Eur J Obstet Gynecol Reprod Biol. 2004; 116(1):3–15. DOI: 10.1016/j.ejogrb.2004.03.007 [PubMed: 15294360]
- 9. Emanuel I. Maternal health during childhood and later reproductive performance. Ann N Y Acad Sci. 1986; 477:27–39. DOI: 10.1111/j.1749-6632.1986.tb40318.x [PubMed: 3545017]
- 10. Ounsted M, Ounsted C. Rate of intra-uterine growth. Nature. 1968; 220(5167):599–600. DOI: 10.1038/220599a0 [PubMed: 5693768]
- 11. Alberman E, Emanuel I, Filakti H, Evans SJW. The contrasting effects of parental birthweight and gestational age on the birthweight of offspring. Paediatr Perinat Epidemiol. 1992; 6:134–44. DOI: 10.1111/j.1365-3016.1992.tb00755.x [PubMed: 1584716]
- 12. Magnus P, Bakketeig LS, Skjaerven R. Correlations of birth weight and gestational age across generations. Annals of Human Biology. 1993; 20(3):231–8. [PubMed: 8489198]
- 13. Conley D, Bennett NG. Race and the Inheritance of Low Birth Weight. Social Biology. 2000; 47(1–2):77–93. DOI: 10.1080/19485565.2000.9989011 [PubMed: 11521458]
- 14. Emanuel I, Kimpo C, Moceri V. The association of maternal growth and socio-economic measures with infant birthweight in four ethnic groups. Int J Epidemiol. 2004; 33(6):1236–42. DOI: 10.1093/ije/dyh269 [PubMed: 15256518]
- 15. Coutinho R, David RJ, Collins JW Jr. Relation of parental birth weights to infant birth weight among African Americans and whites in Illinois: a transgenerational study. Am J Epidemiol. 1997; 146(10):804–9. [PubMed: 9384200]
- 16. Emanuel I, Filakti H, Alberman E, Evans SJ. Intergenerational studies of human birthweight from the 1958 birth cohort. 1. Evidence for a multigenerational effect. Br J Obstet Gynaecol. 1992; 99(1):67–74. DOI: 10.1111/j.1471-0528.1992.tb14396.x [PubMed: 1547177]
- 17. Tavares M, Rodrigues T, Cardoso F, Barros H, Leite LP. Independent effect of maternal birth weight on infant birth weight. J Perinat Med. 1996; 24(4):391–6. [PubMed: 8880637]
- 18. Skjaerven R, Wilcox AJ, Oyen N, Magnus P. Mothers' birth weight and survival of their offspring: population based study. BMJ. 1997; 314(7091):1376–80. DOI: 10.1136/bmj.314.7091.1376 [PubMed: 9161309]
- 19. Shah PS, Shah V, Knowledge Synthesis Group On Determinants Of Preterm/l BWB. Influence of the maternal birth status on offspring: a systematic review and meta-analysis. Acta Obstet Gynecol Scand. 2009; 88(12):1307–18. DOI: 10.3109/00016340903358820 [PubMed: 19916879]
- 20. Hackman E, Emanuel I, van Belle G, Daling J. Maternal birth weight and subsequent pregnancy outcome. JAMA. 1983; 250(15):2016–9. DOI: 10.1001/jama.1983.03340150058027 [PubMed: 6620503]
- 21. Klebanoff MA, Graubard BI, Kessel SS, Berendes HW. Low birth weight across generations. JAMA. 1984; 252(17):2423–7. DOI: 10.1001/jama.1984.03350170025013 [PubMed: 6481929]
- 22. Carr-Hill R, Campbell DM, Hall MH, Meredith A. Is birth weight determined genetically? BMJ (Clin Res Ed). 1987; 295(6600):687–9. DOI: 10.1136/bmj.295.6605.1064-a

- 23. Eriksson J, Forsen T, Tuomilehto J, Osmond C, Barker D. Size at birth, childhood growth and obesity in adult life. Int J Obes Relat Metab Disord. 2001; 25(5):735–40. DOI: 10.1038/sj.ijo. 0801602 [PubMed: 11360158]
- 24. Parsons TJ, Power C, Logan S, Summerbell CD. Childhood predictors of adult obesity: a systematic review. Int J Obes Relat Metab Disord. 1999; 23(Suppl 8):S1–107.
- 25. Frederick IO, Williams MA, Sales AE, Martin DP, Killien M. Pre-pregnancy body mass index, gestational weight gain, and other maternal characteristics in relation to infant birth weight. Maternal and child health journal. 2008; 12(5):557–67. DOI: 10.1007/s10995-007-0276-2 [PubMed: 17713848]
- 26. Enquobahrie DA, Meller M, Rice K, Psaty BM, Siscovick DS, Williams MA. Differential placental gene expression in preeclampsia. Am J Obstet Gynecol. 2008; 199(5):566 e1–11. DOI: 10.1016/ j.ajog.2008.04.020 [PubMed: 18533121]
- 27. American College of Obstetricians and Gynecologists. ACOG Committee opinion no. 548: weight gain during pregnancy. Obstetrics and gynecology. 2013; 121(1):210–2. doi:[http://](http://10.1097/01.AOG.0000425668.87506.4c) [10.1097/01.AOG.0000425668.87506.4c.](http://10.1097/01.AOG.0000425668.87506.4c) [PubMed: 23262962]
- 28. Greenland S. Dose-response and trend analysis in epidemiology: alternatives to categorical analysis. Epidemiology. 1995; 6(4):356–65. [PubMed: 7548341]
- 29. Drake AJ, Walker BR. The intergenerational effects of fetal programming: non-genomic mechanisms for the inheritance of low birth weight and cardiovascular risk. J Endocrinol. 2004; 180(1):1–16. DOI: 10.1677/joe.0.1800001 [PubMed: 14709139]
- 30. Collins J Jr, Rankin K, David R. Low Birth Weight Across Generations: The Effect of Economic Environment. Maternal & Child Health Journal. 2011; 15(4):438–45. DOI: 10.1007/ s10995-010-0603-x [PubMed: 20390329]
- 31. Ncube CN, Enquobahrie DA, Albert SM, Herrick AL, Burke JG. Association of neighborhood context with offspring risk of preterm birth and low birthweight: A systematic review and metaanalysis of population-based studies. Soc Sci Med. 2016; 153:156–64. DOI: 10.1016/j.socscimed. 2016.02.014 [PubMed: 26900890]
- 32. Voldner N, Frey Frøslie K, Godang K, Bollerslev J, Henriksen T. Determinants of birth weight in boys and girls. human\_ontogenetics. 2009; 3(1):7–12. DOI: 10.1002/huon.200900001
- 33. Umbach DM, Wilcox AJ. A technique for measuring epidemiologically useful features of birthweight distributions. Statistics in Medicine. 1996 Jul 15; 15(13):1333–48. [PubMed: 8841645]
- 34. Gilbert JS, Nijland MJ. Sex differences in the developmental origins of hypertension and cardiorenal disease. Am J Physiol Regul Integr Comp Physiol. 2008; 295(6):R1941–52. DOI: 10.1152/ajpregu.90724.2008 [PubMed: 18971349]
- 35. Grigore D, Ojeda NB, Alexander BT. Sex differences in the fetal programming of hypertension. Gender Medicine. 2008; 5:S121–S32. DOI: 10.1016/j.genm.2008.03.012 [PubMed: 18395678]
- 36. Juntunen KS, Kvist AP, Kauppila AJ. A shift from a male to a female majority in newborns with the increasing age of grand grand multiparous women. Hum Reprod. 1997; 12(10):2321–3. DOI: 10.1093/humrep/12.10.2321 [PubMed: 9402303]
- 37. Ricart W, Lopez J, Mozas J, Pericot A, Sancho MA, Gonzalez N, et al. Maternal glucose tolerance status influences the risk of macrosomia in male but not in female fetuses. Journal of epidemiology and community health. 2009; 63(1):64–8. DOI: 10.1136/jech.2008.074542 [PubMed: 18718980]
- 38. Lubchenco LO, Hansman C, Dressler M, Boyd E. Intrauterine Growth as Estimated from Liveborn Birth-Weight Data at 24 to 42 Weeks of Gestation. Pediatrics. 1963; 32:793–800. [PubMed: 14075621]
- 39. Vatten LJ, Skjaerven R. Offspring sex and pregnancy outcome by length of gestation. Early Hum Dev. 2004; 76(1):47–54. DOI: 10.1016/j.earlhumdev.2003.10.006 [PubMed: 14729162]
- 40. Braun JM, Kalkbrenner AE, Calafat AM, Yolton K, Ye X, Dietrich KN, et al. Impact of early-life bisphenol A exposure on behavior and executive function in children. Pediatrics. 2011; 128(5): 873–82. DOI: 10.1542/peds.2011-1335 [PubMed: 22025598]
- 41. Voigt M, Hermanussen M, Wittwer-Backofen U, Fusch C, Hesse V. Sex-specific differences in birth weight due to maternal smoking during pregnancy. Eur J Pediatr. 2006; 165(11):757–61. DOI: 10.1007/s00431-006-0169-1 [PubMed: 16775725]

- 42. Stein AD, Kahn HS, Rundle A, Zybert PA, van der Pal-de Bruin K, Lumey LH. Anthropometric measures in middle age after exposure to famine during gestation: evidence from the Dutch famine. Am J Clin Nutr. 2007; 85(3):869–76. [PubMed: 17344511]
- 43. Tobi EW, Lumey LH, Talens RP, Kremer D, Putter H, Stein AD, et al. DNA methylation differences after exposure to prenatal famine are common and timing- and sex-specific. Hum Mol Genet. 2009; 18(21):4046–53. DOI: 10.1093/hmg/ddp353 [PubMed: 19656776]
- 44. Stark MJ, Clifton VL, Wright IM. Neonates born to mothers with preeclampsia exhibit sex-specific alterations in microvascular function. Pediatr Res. 2009; 65(3):292–5. DOI: 10.1203/PDR. 0b013e318193edf1 [PubMed: 19391250]
- 45. Lampl M, Gotsch F, Kusanovic JP, Gomez R, Nien JK, Frongillo EA, et al. Sex differences in fetal growth responses to maternal height and weight. Am J Hum Biol. 2010; 22(4):431–43. DOI: 10.1002/ajhb.21014 [PubMed: 19950190]
- 46. Rosenfeld CS. Sex-Specific Placental Responses in Fetal Development. Endocrinology. 2015; 156(10):3422–34. DOI: 10.1210/en.2015-1227 [PubMed: 26241064]
- 47. Gabory A, Attig L, Junien C. Sexual dimorphism in environmental epigenetic programming. Mol Cell Endocrinol. 2009; 304(1–2):8–18. DOI: 10.1016/j.mce.2009.02.015 [PubMed: 19433243]
- 48. Gabory A, Roseboom TJ, Moore T, Moore LG, Junien C. Placental contribution to the origins of sexual dimorphism in health and diseases: sex chromosomes and epigenetics. Biol Sex Differ. 2013; 4(1):5.doi: 10.1186/2042-6410-4-5 [PubMed: 23514128]
- 49. Clifton VL. Review: Sex and the human placenta: mediating differential strategies of fetal growth and survival. Placenta. 2010; 31(Suppl):S33–9. DOI: 10.1016/j.placenta.2009.11.010 [PubMed: 20004469]
- 50. Merchant SS, Momin IA, Sewani AA, Zuberi NF. Effect of prepregnancy body mass index and gestational weight gain on birth weight. J Pak Med Assoc. 1999; 49(1):23–5. [PubMed: 10463014]
- 51. Hypponen E, Power C, Smith GD. Parental growth at different life stages and offspring birthweight: an intergenerational cohort study. Paediatr Perinat Epidemiol. 2004; 18(3):168–77. DOI: 10.1111/j.1365-3016.2004.00556.x [PubMed: 15130155]
- 52. Imai K, Keele L, Yamamoto T. Identification, Inference and Sensitivity Analysis for Causal Mediation Effects. Stat Sci. 2010; 25(1):51–71. DOI: 10.1214/10-Sts321
- 53. Nieto-Garcia FJ, Bush TL, Keyl PM. Body mass definitions of obesity: sensitivity and specificity using self-reported weight and height. Epidemiology. 1990; 1(2):146–52. [PubMed: 2073502]
- 54. Sanderson M, Williams MA, White E, Daling JR, Holt VL, Malone KE, Self SG, Moore DE. Validity and reliability of subject and mother reporting of perinatal factors. American journal of epidemiology. 1998 Jan 15; 147(2):136–40. [PubMed: 9457002]
- 55. Tehranifar P, Liao Y, Flom JD, Terry MB. Validity of self-reported birth weight by adult women: sociodemographic influences and implications for life-course studies. American journal of epidemiology. 2009:kwp205.
- 56. Emanuel I, Leisenring W, Williams MA, Kimpo C, Estee S, O'Brien W, et al. The Washington State Intergenerational Study of Birth Outcomes: methodology and some comparisons of maternal birthweight and infant birthweight and gestation in four ethnic groups. Paediatr Perinat Epidemiol. 1999; 13(3):352–69. DOI: 10.1046/j.1365-3016.1999.00184.x [PubMed: 10440054]



#### **Figure. Sex-specific associations of maternal and offspring birthweight**

Fitted values. Model adjusted for potential confounding variables: maternal race, family history of diabetes, maternal smoking history and educational attainment; precision variables: age (as a linear spline), marital status, and parity

#### **Table 1**

Selected Study Participant Characteristics (N=3736)



*Note:* BMI = body mass index. Underweight (<18.5 kg/m<sup>2</sup>), normal weight (18.5–24.99 kg/m<sup>2</sup>) and overweight/obese (25 kg/m<sup>2</sup>).

Weight gain during pregnancy based on Institute of Medicine recommendations per BMI category

#### **Table 2**

Associations of maternal birthweight with offspring birthweight and offspring risk of low birthweight



*Note*:  $BW = birthweight$ .  $LBW = low birthweight$ .

\* p-value < 0.05;

\*\*

p-value < 0.0001; Adjusted  $R^2$  = variation in offspring BW explained

 ${}^{a}$ Model 1 - Unadjusted: crude change in mean infant BW.

 $b$  Model 2 - adjusted: adjusted for potential confounding variables: maternal race, family history of diabetes, maternal smoking history and educational attainment; precision variables: age (as a linear spline), marital status, parity, and offspring-sex.

 $c$ Model 3 –adjusted: adjusted for Model 2 variables plus gestational diabetes, preeclampsia, chronic hypertension, pre-pregnancy body mass index, and preterm birth.

d Per 100 grams maternal birthweight.

 $e$ Comparing LBW mothers and non-LBW mothers (reference).

#### **Table 3**

Offspring sex-specific associations of maternal birthweight with offspring birthweight and offspring risk of low birthweight



*Note*:  $BW = birthweight$ .  $LBW = low birthweight$ .

\*

# **Author Manuscript** Author Manuscript

p-value < 0.05;

\*\*p-value < 0.0001; Adjusted  $R^2$  = variation in offspring BW explained

 $a$  Model adjusted for potential confounding variables: maternal race, family history of diabetes, maternal smoking history and educational attainment; precision variables: age (as a linear spline), marital status, parity, and offspring-sex.

 $b$ Per 100 grams maternal birthweight.

 $c_{\text{Comparing LBW}$  mothers and non-LBW mothers (reference).

#### **Table 4**

Associations of maternal birthweight with offspring birthweight and offspring risk of low birthweight, by maternal pre-pregnancy body mass index



*Note*: BW = birthweight. LBW = low birthweight. BMI = body mass index.

\* p-value < 0.05;

\*\* p-value < 0.0001

 ${}^{a}$ Model adjusted for potential confounding variables: maternal race, family history of diabetes, maternal smoking history and educational attainment; precision variables: age (as a linear spline), marital status, parity, and offspring-sex.

b<br>Per 100 grams maternal birthweight.

 $c$ Comparing LBW mothers and non-LBW mothers (reference).